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The Journal of Political Economy, Vol. 98, No. 3. (Jun., 1990), pp. 519-543.

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Habit Formation: A Resolution of the Equity Premium Puzzle

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The equity premium puzzle, identified by Mehra and Prescott, states that, for plausible values of the risk aversion coefficient, the difference of the expected rate of return on the stock market and the riskless rate of interest is too large, given the observed small variance of the growth rate in per capita consumption. The puzzle is resolved in the context of an economy with rational expectations once the time separability of von Neumann–Morgenstern preferences is relaxed to allow for adjacent complementarity in consumption, a property known as habit persistence. Essentially habit persistence drives a wedge between the relative risk aversion of the representative agent and the intertemporal elasticity of substitution in consumption.

I. Introduction

Rational expectations, a cornerstone of modern theories in economics and finance, has come under attack. Are prices too volatile relative to the information arriving in the market? Is the mean premium on equities over the riskless rate too large? Is the real interest rate too low? Is the market's risk aversion too high? Is the intertemporal elasticity of substitution in consumption with respect to changes in the

Partial support for this project was provided by a National Science Foundation grant and by a gift to the Graduate School of Business, University of Chicago, by Dimensional Fund Advisors. I am grateful to Gary Becker, Fischer Black, Stephen Cecchetti, John Cochrane, Martin Eichenbaum, Wayne Ferson, Lars Hansen, Robert Lucas, Rajnish Mehra, Kevin M. Murphy, Dan Nelson, Ed Prescott, Ken Singleton, and Arnold Zellner for helpful comments. Expert computational assistance was provided by Ed Allen and Andrew Karolyi. I remain responsible for errors.

[Journal of Political Economy, 1990, vol. 98, no. 3]

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productivity of capital too low? Finally, is the time series of aggregate consumption of nondurables and services too smooth?

Mehra and Prescott (1985) raised some of these questions in their equity premium puzzle. They employed a variant of Lucas's (1978) pure exchange economy and conducted a "calibration" exercise in the spirit of Kydland and Prescott (1982). Mehra and Prescott chose the parameters of the endowment process to match the sample mean, variance, and first-order autocorrelation of the annual growth rate of per capita consumption in the years 1889–1978. They postulated that the representative agent has time- and state-separable utility. The puzzle is that they were unable to find a plausible pair of the subjective discount rate and relative risk aversion (RRA) of the representative agent to match the sample mean of the annual real rate of interest and of the equity premium over the same 90-year period. Stated differently, the consumption growth rate appears to be too smooth to justify the mean equity premium.

The equity premium puzzle is not an isolated observation. Hansen and Singleton (1982, 1983), Ferson (1983), Grossman, Melino, and Shiller (1987), and several others rejected the Euler equation restriction on asset returns and the marginal rate of substitution implied by time- and state-separable preferences. Mankiw, Romer, and Shapiro (1985), Campbell and Shiller (1988), and West (1988*a*, 1988*b*) found that stock returns are too volatile if future dividends are discounted at a constant rate. Some of the empirical literature on the consumption function rejected the joint hypothesis of rationality and time- and state-separable preferences (see Deaton 1987; Hall 1989). Black (1986) and Roll (1988) questioned the rationality of price changes. Campbell and Kyle (1988) and De Long et al. (1990) made a case for noise traders. Finally, the stock market crash of October 1987 has added fuel to the debate.

The goal of this paper is to show that the equity premium puzzle is resolved in a rational expectations model, once we relax the time separability of preferences and allow for adjacent complementarity in consumption, a property known as habit persistence.

Marshall (1920) discussed the notion that tastes can be cultivated and that they are affected by past consumption. Duesenberry's (1949) thesis on the consumption function is probably the first serious examination of the implications of habit persistence. Ryder and Heal (1973) introduced the notion of adjacent and distant complementarity and discussed the stability of a growth model in the presence of habit persistence. Stigler and Becker (1977) argued that preferences should not be taken as exogenous but that it is fruitful to endogenize them and search for factors that explain differences or changes in behavior. Kydland and Prescott (1982) introduced preferences that are non-

time separable in leisure. Becker and Murphy (1988) presented a theory of rational addiction and provided an insightful discussion on the link between addiction and complementarity. Sundaresan (1989) discussed the volatility of consumption and wealth in the presence of habit persistence.

In his examination of habit formation and dynamic demand functions, Pollak (1970, p. 761) insisted that “a fundamental assumption of the habit-formation model is that the individual does not take account of the effect of his current purchase on his future preferences and future consumption.” I see nothing fundamental in the association of habit formation with some form of myopia or irrationality. In the present paper, habit persistence is introduced in a model with rational expectations, given that the goal is to show that the equity premium puzzle does not lead to the conclusion that the rational expectations model is bankrupt.

The paper is organized as follows. Habit persistence is embedded in a variant of the neoclassical growth model in Section II. Theorem 1 proves existence and uniqueness of an optimal policy and presents the optimal policy, the derived utility of capital, and the dynamics of capital and consumption. Theorem 2 derives the stationary distribution of the state variable and enables one to calculate the unconditional mean and variance of the consumption growth rate. Section IIC illustrates that the key role of habit persistence is to drive a wedge between the RRA coefficient and the inverse of the intertemporal elasticity of substitution in consumption. In Section III, I interpret the growth model as the equilibrium in a representative-consumer production economy. I resolve the equity premium puzzle by showing in table 1 that habit persistence can generate the sample mean and variance of the consumption growth rate with low risk aversion. In Section IV, I discuss alternative potential explanations of the puzzle. Finally, in Section V, I review related empirical evidence and offer suggestions for future research.

II. Habit Persistence in a Production Economy

A. *The Model and Assumptions*

Habit persistence is introduced in a variant of the neoclassical growth model. The optimal consumption and investment paths are interpreted as the equilibrium paths in a representative-consumer production economy, and the shadow prices of assets are interpreted as the equilibrium prices.

There exists only one production good, which is also the consumption good. This good may be consumed or invested in two tech-

nologies. The technologies have constant returns to scale and rates of return over the period $[t, t + dt]$ equal to $r dt$ and $\mu dt + \sigma dw(t)$, respectively, where r , μ , and σ are constants and $w(t)$ is a standard Brownian motion in R^1 .

The infinitely lived representative consumer has capital $W(t)$ at time t denominated in units of the consumption good, investing fraction $\alpha(t)$, $0 \leq \alpha(t) \leq 1$, of the capital in the risky technology and the remaining fraction $1 - \alpha(t)$ in the riskless technology. The consumer also consumes $c(t)dt$ in the period $[t, t + dt]$. Assume zero endowment flow and labor income. The increase in capital over $[t, t + dt]$ is

$$dW(t) = \{[(\mu - r)\alpha(t) + r]W(t) - c(t)\}dt + \sigma\alpha(t)W(t)dw(t). \quad (1)$$

Given a consumption and investment policy, $\{c(t), \alpha(t), t \geq 0\}$, the expected utility of consumption from time 0 to infinity is defined as

$$E_0 \int_0^\infty e^{-\rho t} \gamma^{-1} [c(t) - x(t)]^\gamma dt, \quad (2)$$

where

$$x(t) \equiv e^{-at} x_0 + b \int_0^t e^{a(s-t)} c(s) ds. \quad (3)$$

Since $\lim_{\gamma \rightarrow 0} \gamma^{-1}(y^\gamma - 1) = \ln y$, the case $\gamma \rightarrow 0$ in equation (2) corresponds to logarithmic utility, which may be treated separately rather than cluttering the notation by replacing $(c - x)^\gamma$ with $(c - x)^\gamma - 1$.

The special case $x_0 = b = 0$ corresponds to time-separable utility with constant RRA, $1 - \gamma$. The novel feature of the utility function studied in this paper is that the subsistence level of consumption, $x(t)$, is an exponentially weighted sum of past consumption. Thus utility is not time separable but exhibits habit persistence. The particular form of the habit-forming state variable, $x(t)$, defined in equation (3), was introduced by Ryder and Heal (1973), who studied a two-factor growth model maximizing expected utility $E_0 \int_0^\infty e^{-\rho t} u(c(t), x(t)) dt$.

The utility function defined in equations (2) and (3) exhibits adjacent complementarity in consumption; that is, an increase in consumption increases the marginal utility of consumption at adjacent dates relative to the marginal utility of consumption at distant ones. Formally, define $J'(c(\cdot), t_1)$ as the marginal utility of consumption at date t_1 , where the derivative takes into account the impact of the change in $c(t_1)$ on all future values of $x(t)$. Define also the marginal rate of substitution between consumption at dates t_1 and t_2 , $0 < t_1 < t_2$, as $J'(c(\cdot), t_1)/J'(c(\cdot), t_2)$. By specializing the results of Ryder and Heal (1973), one can show that along a constant consumption path, $c(t) =$

$ax(t)/b = ax_0/b$, there exists a number \bar{t} , $t_1 < \bar{t} < t_2$, such that the marginal rate of substitution increases when consumption $c(t)$ increases, for $t < \bar{t}$.

Theoretical and empirical tractability is the only reason why I model habit formation as in equations (2) and (3). My goal is not to study the most general utility function that exhibits habit persistence but rather to employ the simplest utility specification that resolves the equity premium puzzle.

The consumer's choice of a consumption and investment policy is restricted to the set of admissible policies defined by the following four properties: (i) The consumption and investment decisions taken at date t are based solely on information available at date t . (ii) The consumption rate is nonnegative ($c(t) \geq 0$), does not fall below the subsistence level ($c(t) \geq x(t)$), and results in finite total consumption over any horizon; that is, $\int_0^t c(s)ds < \infty$, for all t almost surely. (iii) Investment in both technologies is nonnegative; that is, $0 \leq \alpha(t) \leq 1$ for all t almost surely. (iv) The policy guarantees that the capital remains nonnegative; that is, $W(t) \geq 0$ for all t almost surely.

An optimal admissible policy and the associated derived utility of capital are defined by

$$V(W_0, x_0) \equiv \max_{\substack{\text{admissible} \\ \{\alpha(s), c(s), s \geq 0\}}} E_0 \int_0^\infty e^{-\rho s} \gamma^{-1} [c(s) - x(s)]^\gamma ds, \tag{4}$$

where $W(0) = W_0$ and $x(0) = x_0$. I impose restrictions on the model parameters and motivate these restrictions.

Assume that

$$1 - \gamma > 0, \quad \gamma \neq 0. \tag{5}$$

The case $\gamma = 0$ corresponds to logarithmic utility and may be treated separately. Condition (5) is necessary if the RRA coefficient of the consumer is to be positive. As shown later, $1 - \gamma$ is only approximately equal to the RRA coefficient. The equality is exact if utility is time separable, $b = 0$, and is of the power form $x_0 = 0$.

Conditions (6)–(8),

$$W_0 > 0, \tag{6}$$

$$W_0 - \frac{x_0}{r + a - b} > 0, \tag{7}$$

and

$$0 < b < r + a, \tag{8}$$

jointly imply that the set of admissible policies is nonempty. In particular, the policy $\{c(t) = (r + a - b)W(t), \alpha(t) = 0, t \geq 0\}$ implies

$$W(t) = W_0 \exp[(b - a)t] > 0,$$

$$c(t) = (r + a - b)W_0 \exp[(b - a)t] > 0,$$

$$\int_0^t c(s)ds < \infty,$$

$$c(t) - x(t) = (r + a - b) \left(W_0 - \frac{x_0}{r + a - b} \right) \exp(-at) > 0,$$

thereby satisfying all the conditions of an admissible policy.

The condition

$$\rho - \gamma r - \frac{\gamma(\mu - r)^2}{2(1 - \gamma)\sigma^2} > 0 \quad (9)$$

ensures that, under the optimal policy, the expected utility of consumption flow grows at a rate that is lower than the time preference, so that the expected utility of consumption over the infinite horizon is finite. It also implies that the appropriate transversality condition is satisfied.

Finally, assume that

$$x_0 \geq 0 \quad (10)$$

and

$$0 \leq m \leq 1, \quad m \equiv \frac{\mu - r}{(1 - \gamma)\sigma^2}. \quad (11)$$

Conditions (5)–(11) are invoked in theorem 1 to prove that an optimal policy exists and is unique, and they lead to closed-form expressions for the optimal policy and for the derived utility of capital. Essentially, condition (10) guarantees that the condition $c(t) \geq 0$ of an admissible policy is nonbinding, and condition (11) guarantees that the condition $0 \leq \alpha(t) \leq 1$ of an admissible policy is nonbinding. Then the optimal consumption and investment are at an interior maximum, and this simplification leads to closed-form expressions.

B. Optimal Consumption and Investment Policy

In the first theorem I prove existence and uniqueness of an optimal policy, state the optimal policy, state the derived utility of capital, and state the dynamics of capital and consumption.

THEOREM 1. Under conditions (5)–(11), an optimal admissible consumption and investment policy exists, is unique, and is given by

$$c^*(t) = x(t) + h \left[W(t) - \frac{x(t)}{r + a - b} \right] \tag{12}$$

and

$$\alpha^*(t) = m \left[1 - \frac{x(t)/W(t)}{r + a - b} \right], \tag{13}$$

where

$$h \equiv \left[\frac{r + a - b}{(r + a)(1 - \gamma)} \right] \left[\rho - \gamma r - \frac{\gamma(\mu - r)^2}{2(1 - \gamma)\sigma^2} \right] > 0. \tag{14}$$

The derived utility of capital is

$$V(W(t), x(t)) = \frac{(r + a - b)h^{\gamma-1}}{(r + a)\gamma} \left[W(t) - \frac{x(t)}{r + a - b} \right]^{\gamma}. \tag{15}$$

The capital is

$$W(t) = \frac{x(t)}{r + a - b} + \left(W_0 - \frac{x_0}{r + a - b} \right) \times \exp \left[\left(n - \frac{m^2\sigma^2}{2} \right) t + m\sigma w(t) \right], \tag{16}$$

and the consumption growth rate is

$$\frac{dc(t)}{c(t)} = \left[n + b - \frac{(n + a)x(t)}{c(t)} \right] dt + \left[1 - \frac{x(t)}{c(t)} \right] m\sigma dw(t), \tag{17}$$

where

$$n \equiv \frac{r - \rho}{1 - \gamma} + \frac{(\mu - r)^2(2 - \gamma)}{2(1 - \gamma)^2\sigma^2}. \tag{18}$$

The theorem is proved in Appendix A. Merton (1971) considered the special case $a = b = 0$, which corresponds to time-separable utility with hyperbolic absolute risk aversion. He stated the optimal policy and proved its optimality and uniqueness. Sundaresan (1989) stated the optimal policy in two cases of nonseparable utility, $a = b$ and $a \geq b$, but the direct utility exponential is $c - x$.

In addressing the equity premium puzzle, I interpret the optimal paths specified by theorem 1 as the equilibrium paths in a representative-consumer economy. In particular, the consumption growth rate, specified by equation (17), is interpreted as the per capita consump-

tion growth rate. The mean and variance of the consumption growth rate are functions of the state variable $x(t)$, which appears as the ratio $z(t) \equiv x(t)/c(t)$. Theorem 2 states conditions under which this ratio has a stationary distribution and presents this distribution. This distribution is used to calculate the unconditional mean and variance of the consumption growth rate.

The RRA coefficient is defined in Section IIC and is shown to be a function of the state variable $x(t)$, which appears as the ratio $y(t) \equiv x(t)/[c(t) - x(t)] = z(t)/[1 - z(t)]$. Theorem 2 states conditions under which this ratio has a stationary distribution and presents this distribution and the mean of $y(t)$. This distribution is used to calculate the unconditional mean of the RRA coefficient.

THEOREM 2. Assume that conditions (5)–(11) hold and also that

$$n + a - b - m^2\sigma^2 > 0. \tag{19}$$

Then (i) $y(t) \equiv x(t)/[c(t) - x(t)]$ has a stationary probability distribution with density

$$p_y(y) = ky^{-2(n+a-b)/m^2\sigma^2} e^{-2b/m^2\sigma^2 y}, \quad 0 \leq y < \infty, \tag{20}$$

where

$$k^{-1} \equiv \left(\frac{2b}{m^2\sigma^2}\right)^{1-2(n+a-b)/m^2\sigma^2} \Gamma\left[\frac{2(n+a-b)}{m^2\sigma^2} - 1\right] \tag{21}$$

and $\Gamma(\cdot)$ is the gamma function. For the stationary distribution, y has a single mode \hat{y} ,

$$\hat{y} = \frac{b}{n + a - b} < \infty, \tag{22}$$

and mean

$$\bar{y} = \frac{b}{n + a - b - m^2\sigma^2} < \infty. \tag{23}$$

(ii) $z(t) \equiv x(t)/c(t)$ has a stationary probability distribution with density

$$p_z(z) = ke^{2b/m^2\sigma^2} (1 - z)^{2(n+a-b-m^2\sigma^2)/m^2\sigma^2} z^{2(b-a-n)/m^2\sigma^2} e^{-2b/m^2\sigma^2 z}, \tag{24}$$

$$0 \leq z < 1.$$

For the stationary distribution, z has a single mode \hat{z} ,

$$\hat{z} = \frac{n + a - [(n + a)^2 - 4m^2\sigma^2 b]^{1/2}}{2m^2\sigma^2}. \tag{25}$$

The proof is given in Appendix B. Using equation (17), we can calculate the unconditional mean and variance of the consumption growth rate as

$$\frac{E(dc/c)}{dt} = n + b - (n + a) \int_0^1 zp_z(z)dz \tag{26}$$

and

$$\frac{\text{var}(dc/c)}{dt} = m^2\sigma^2 \int_0^1 (1 - z)^2 p_z(z)dz. \tag{27}$$

The density $p_z(z)$ is given in theorem 2, and the integration is done numerically since we are unable to obtain closed-form expressions for the integrals.

C. A Wedge between the RRA Coefficient and the Inverse of the Intertemporal Elasticity of Substitution in Consumption

I define the RRA coefficient and the intertemporal elasticity of substitution in consumption (s). I show that the product $s \cdot \text{RRA}$ equals one in the time-separable model ($b = 0$) but is substantially below one in the nonseparable model and for the particular parameter values that resolve the equity premium puzzle. Thus habit persistence drives a wedge between the RRA coefficient and the inverse of the intertemporal elasticity of substitution in consumption.

I define the RRA coefficient in terms of an atemporal gamble that changes the current level of capital by the outcome of the gamble and is given by

$$\text{RRA} = \frac{-WV_{ww}}{V_w} = \frac{1 - \gamma}{1 - \{x(t)/[W(t)(r + a - b)]\}}. \tag{28}$$

This definition is consistent with that in Giovannini and Weil (1988) for Kreps-Porteus preferences. In the context of an intertemporal model it would be improper to define the RRA coefficient in terms of an atemporal gamble that changes either current consumption or consumption at some specified future date by the outcome of the gamble.

The RRA coefficient is a function of wealth and of the state variable $x(t)$. A sudden drop in wealth leaves $x(t)$ unchanged in the short run and increases the RRA coefficient. This drop is only temporary because the RRA coefficient has a stationary distribution. To see this, we can use equation (12) to eliminate $W(t)$ from equation (28) and obtain

$$\text{RRA} = (1 - \gamma) \left[1 + \frac{hy(t)}{r + a - b} \right]. \tag{29}$$

The RRA coefficient has a steady-state distribution because $y(t)$ does, by theorem 2. Since the mean value of $y(t)$ in the steady state is given by equation (23), we obtain the mean of the RRA coefficient as

$$\overline{\text{RRA}} = (1 - \gamma) \left[1 + \frac{hb}{(r + a - b)(n + a - b - m^2\sigma^2)} \right]. \quad (30)$$

With the parameter values that resolve the equity premium puzzle, I show in Section III that the mean of the RRA coefficient is of the same order of magnitude as $1 - \gamma$.

The elasticity of substitution in consumption is defined here as the derivative of the expected growth rate in consumption with respect to r , with $z(t)$, $\mu - r$, and σ^2 held constant:

$$s \equiv \frac{\partial[E(dc/c)/dt]}{\partial r} \Big|_{z(t), \mu - r, \sigma^2} = \frac{1 - z(t)}{1 - \gamma}. \quad (31)$$

Note that the elasticity may also be defined as the inverse of the expression $-cu_{cc}/u_c$. I stress, however, that this expression need not equal the RRA coefficient because risk aversion is defined in terms of an atemporal gamble that changes wealth and not in terms of a gamble that changes consumption.

We can combine equations (29) and (31) and write the product of the elasticity of substitution and the RRA coefficient as

$$\begin{aligned} s \cdot \text{RRA} &= (1 - z) \left(1 + \frac{hy}{r + a - b} \right) \\ &= 1 - \left(1 - \frac{h}{r + a - b} \right) z. \end{aligned} \quad (32)$$

To consider the special case of time-separable utility, let $b \rightarrow 0$. By equation (25) the modal value of z tends to zero, and therefore the modal value of the product $s \cdot \text{RRA}$ tends to one. Note that we do not assume that $x_0 = 0$ or $a > 0$; therefore $x(t)$ need not vanish as $b \rightarrow 0$. It is the assumption that utility is time separable, and not the stronger assumption that $x(t)$ vanishes, that gives the result that the product $s \cdot \text{RRA}$ has modal value one.

In his insightful exposition of growth theory, Solow (1970, p. 85) proved in the context of a deterministic growth model with time-separable utility that the consumption growth rate is linear in the net marginal product of capital, with the coefficient equal to the inverse of the RRA coefficient. Put differently, the intertemporal elasticity of substitution in consumption is the inverse of the RRA coefficient. The assumption that consumption growth is deterministic is not crucial. Hansen and Singleton (1983), Breeden (1986), and Hall (1988) extended the result under uncertainty by making reasonable assump-

tions about the stochastic process of consumption and the rates of return.

With habit persistence, the mode of the stationary distribution of z is given by equation (25). In the next section I show that the modal value of $s \cdot \text{RRA}$ is substantially below one at the parameter values that resolve the equity premium puzzle.

Hansen and Singleton (1982, 1983), Ferson (1983), Grossman et al. (1987), and others rejected the Euler equation restriction implied by the time-separable model. Hall (1988) argued that since the time-separable model forces the product $s \cdot \text{RRA}$ to equal one, these results are rejections of the Euler equation and the hypothesis that the product $s \cdot \text{RRA}$ equals one. Ferson and Constantinides (1989) rejected the Euler equation implied by the time-separable model when the alternative hypothesis is the Euler equation implied by the non-separable model. These results may be interpreted as evidence against the restriction $s \cdot \text{RRA} = 1$.

For the equilibrium of the particular model developed in this section, we may write the capital elasticity of consumption as

$$\left. \frac{\partial c/c}{\partial W/W} \right|_x = s \cdot \text{RRA} \quad (33)$$

and the ratio of the standard deviation of the consumption growth rate and the capital growth rate as

$$\frac{\text{std}(dc/c)}{\text{std}(dW/W)} = s \cdot \text{RRA}. \quad (34)$$

Since habit persistence allows the product $s \cdot \text{RRA}$ to be substantially below one, we can conclude that habit persistence smooths consumption growth over and above the smoothing implied by the life cycle-permanent income hypothesis with time-separable utility.

III. Resolution of the Equity Premium Puzzle

I interpret the growth model developed in Section II as the equilibrium in a representative-consumer production economy. The optimal consumption path is interpreted as the per capita consumption.

Mehra and Prescott (1985) estimated the mean of the annual growth rate of per capita real consumption of nondurables and services in the years 1889–1978 to be .0183 with a range $-.0025, .03$ in subperiods. They also estimated the standard deviation of the growth rate in the years 1889–1978 to be .0357 with a range $.010, .0528$ in subperiods. In terms of our notation, we want the model stated in Section II to imply $E(dc/c)/dt = .0183$ per year and $\text{var}(dc/c)/dt = (.0357)^2$ per year.

Mehra and Prescott estimated the mean annual real rate of return on a relatively riskless security to be .008, using 90-day Treasury bills in the 1931–78 period, Treasury certificates in the 1920–30 period, and 60–90-day prime commercial paper in the 1889–1920 period. Thus we can set $r = .01$ per year.

Let us introduce a firm that has capital $K(t)$ at time t . The firm has free access to the two production technologies. It invests capital $\delta_1 K(t)$ in the risky technology and the remaining capital $(1 - \delta_1)K(t)$ in the riskless technology, where δ_1 is a constant, $0 < \delta_1 \leq 1$. The firm is financed with equity of value $S(t)$ and riskless debt of value $B(t)$. The firm maintains the ratio $S(t)/[S(t) + B(t)] = \delta_2$ constant, $0 < \delta_2 \leq 1$. Since the firm has free access to the constant-returns-to-scale technologies, the value of the firm equals its capital, that is, $S(t) + B(t) = K(t)$. Since the bonds are riskless, their rate of return is $dB/B = rdt$. Denoting by dS/S the rate of return on equity, we obtain

$$dS(t) + B(t)rdt = \delta_1 K(t)[\mu dt + \sigma dw(t)] + (1 - \delta_1)K(t)rdt,$$

which simplifies into

$$\frac{dS(t)}{S(t)} = \left(\frac{\delta_1}{\delta_2}\right)[(\mu - r)dt + \sigma dw(t)] + rdt. \quad (35)$$

I interpret the equity of the firm as a portfolio of the stocks represented in the Standard and Poor's composite stock price index. Given the leverage (δ_2) of the firms represented in the index, the ratio δ_1/δ_2 is free in the range $0 < \delta_1/\delta_2 \leq \delta_2^{-1}$ since the parameter δ_1 is free in the range $0 < \delta_1 \leq 1$. In our calculation, we can set $\delta_1/\delta_2 = 1$, which is consistent with any amount of leverage.

Mehra and Prescott estimated the annual real return on the Standard and Poor's composite stock price index in the 1889–1978 period to have mean .0698 (with range $-.0014, .1896$ in subperiods) and standard deviation .1654 (with range $.002, .2790$ in subperiods). These estimates are generally consistent with those by Ibbotson and Sinquefeld (1982, p. 15). Thus we can set

$$\frac{E(dS/S)}{dt} = \left(\frac{\delta_1}{\delta_2}\right)(\mu - r) = .06 \text{ per year} \quad (36)$$

and

$$\frac{\text{var}(dS/S)}{dt} = \left(\frac{\delta_1}{\delta_2}\right)^2 \sigma^2 = (.165)^2 \text{ per year.} \quad (37)$$

The mean and variance of the consumption growth rate are independent of the ratio δ_1/δ_2 . Equations (26) and (27) show that the mean and variance of the consumption growth rate depend on the parameters μ and σ only in the combination $(\mu - r)/\sigma = .06/.165$, which is

independent of δ_1/δ_2 . However, condition (11) requires $1 - \gamma \geq 2.2(\delta_1/\delta_2)$.

In the context of the production economy, time-separable utility implies that the RRA coefficient equals 10.2; hence the equity premium puzzle. To see this, time-separable utility implies that $b \rightarrow 0$ and that the modal value of $z(t)$ is zero. Then

$$\frac{\text{var}(dc/c)}{dt} = m^2\sigma^2 = \frac{(\mu - r)^2}{\sigma^2(1 - \gamma)^2} = (.0357)^2,$$

which implies $1 - \gamma = 10.2$, irrespective of the ratio δ_1/δ_2 .¹

I proceed to show that habit persistence can generate the sample mean and variance of the consumption growth rate with a low RRA coefficient. Let us set $\rho = .037$ per year, $1 - \gamma = 2.2$, and $\delta_1/\delta_2 = 1$. The reader may verify that these parameter values, together with the parameter values specified in equations (36) and (37), satisfy the model conditions (5), (9), and (11).

Let us consider pairs of parameter values (a , b) that satisfy the conditions (8) and (19). For each pair (a , b), the stationary distribution of z is given by equation (24). We can calculate the mean and variance of the consumption growth rate by performing the numerical integration in equations (26) and (27). Table 1 reports pairs (a , b) for which the mean and variance of the consumption growth rate match their sample estimates.

The table also reports the mean RRA coefficient. As one shifts to the right of the table, the mean RRA coefficient decreases and approaches the value $1 - \gamma = 2.2$. The equity premium puzzle is resolved in the sense that the model generates the mean and variance of the consumption growth rate with the mean RRA coefficient as low as 2.81.

If the value of 2.81 is not sufficiently low relative to the reader's prior on the RRA coefficient, we can generate the target mean and variance of the consumption growth rate with a lower RRA coefficient by setting a lower value for $1 - \gamma$. Now in order to satisfy the condition (11), we have to set $\delta_1/\delta_2 > 1$. If we assume that the firm has a debt/equity ratio equal to one, then $\delta_2 = .5$ and we can set $\delta_1/\delta_2 = 2$

¹ Friend and Blume (1975) estimated the demand for risky assets and inferred the RRA coefficient to be well below 10, under the assumption that the investment opportunity set is constant. Black (1988) and Kocherlakota (1988) pointed out that the Friend and Blume inference of the RRA coefficient is invalid if the investment opportunity set is not constant. An alternative source of estimates of the RRA coefficient, which does not rely on the assumption of a constant investment opportunity set, is based on the Euler equation implied by time-separable and non-time-separable utility functions. Typically, risk aversion is estimated to be well below 10. Some of this literature is reviewed in Sec. V.

TABLE I
MEAN AND VARIANCE OF THE CONSUMPTION GROWTH RATE GENERATED
BY THE MODEL WITH HABIT PERSISTENCE

Parameter a , per year	.1	.2	.3	.4	.5	.6
Parameter b	.093	.172	.250	.328	.405	.492
Mode (\hat{z}) of the state variable z	.86	.82	.81	.80	.79	.81
Mean annual growth rate in consumption:						
Unconditional mean	.018	.019	.018	.018	.018	.018
At $z = \hat{z}$.011	.013	.014	.014	.014	.014
Standard deviation of the annual growth rate in consumption:						
Unconditional mean	.036	.036	.036	.036	.036	.034
At $z = \hat{z}$.023	.029	.032	.033	.034	.032
RRA coefficient:						
Unconditional mean	8.67	4.37	3.47	3.09	2.88	2.81
At $z = \hat{z}$	7.03	4.09	3.36	3.03	2.84	2.78
Elasticity of substitution (s)						
at $z = \hat{z}$.06	.08	.09	.09	.09	.09
$s \cdot \text{RRA at } z = \hat{z}$.42	.33	.30	.27	.26	.25

NOTE.—The assumed parameter values are $r = .01$, the annual rate of return of the riskless technology; $\mu - r = .06$, the difference between the mean annual rate of return of the risky technology and the annual rate of return of the riskless technology; $\sigma = .165$, the standard deviation of the annual rate of return of the risky technology; $\gamma = -1.2$, the power in the utility function; and $\rho = .037$, the rate of time preference in units (year)⁻¹.

and $1 - \gamma = 1.1$. By judicious choice of parameters (a , b), we can generate the target mean and variance of the consumption growth rate with the RRA coefficient close to 1.1.

An interesting feature of table 1 is that the modal value of the state variable $z(t) = x(t)/c(t)$ is about .8 for all the reported (a , b) pairs. The model predicts that the subsistence level of consumption, $x(t)$, generated by habit persistence, is about 80 percent of the level of consumption. This prediction is discussed further in the last section. Another interesting feature of the table is that the intertemporal elasticity of substitution in consumption is substantially below one. Finally, the product of the elasticity of substitution and the RRA coefficient is about .25 for the pairs (a , b) that resolve the equity premium puzzle with a low RRA coefficient. This illustrates the key role of habit persistence in resolving the puzzle by driving a wedge between the RRA coefficient and the inverse of the elasticity of substitution.

IV. Discussion

We have resolved the equity premium puzzle by relaxing Mehra and Prescott's (1985) assumption that utility is time separable. However,

our economy differs from theirs in two other respects as well. First, our economy allows for production while theirs is an exchange economy. Second, theirs is a discrete-time economy in which the state is a Markov process with two realizations, while ours is a continuous-time economy in which the forcing process is a diffusion. To make the case that habit persistence is the key to the puzzle, we need to demonstrate that these two differences in modeling the economy are inessential.

The first difference is inessential because, as Mehra and Prescott (1985) and Mehra (1988) pointed out, the task of explaining the puzzle in a production economy is not easier than in an exchange economy. The introduction of production does not increase the set of joint equilibrium processes on consumption and asset prices. In fact it may be harder to explain the puzzle in a production economy because the consumption process is no longer exogenous but must be obtained as the equilibrium outcome.

The second difference is inessential as well. I demonstrate that time separability in preferences is the key restriction that generates the puzzle in Mehra and Prescott's economy. Let m_{t+1} be the marginal rate of substitution and R_{Ft} be the one-plus riskless rate of interest between periods t and $t + 1$. The Euler equation states that

$$E(m_{t+1}R_{Ft}|I_t) = 1, \quad (38)$$

where I_t is the public information in period t . Since R_{Ft} is in the information set I_t , we can write $E(m_{t+1}|I_t) = R_{Ft}^{-1}$ and, by Jensen's inequality, express the unconditional mean of the marginal rate of substitution as

$$E(m) = E(R_F^{-1}) \geq [E(R_F)]^{-1}. \quad (39)$$

Let equity have one-plus rate of return R_{t+1} . The Euler equation states that

$$E(m_{t+1}R_{t+1}|I_t) = 1 \quad (40)$$

and

$$E(mR) = 1. \quad (41)$$

Following the methodology of Hansen and Jagannathan (1988), we can write

$$\begin{aligned} 1 &= E(mR) = E(m)E(R) + \text{cov}(m, R) \\ &\geq E(m)E(R) - \text{std}(m) \text{std}(R) \\ &\geq \frac{E(R)}{E(R_F)} - \text{std}(m) \text{std}(R) \end{aligned}$$

by equation (39) or

$$\text{std}(m) \geq \frac{[E(R)/E(R_F)] - 1}{\text{std}(R)}. \quad (42)$$

Assume that utility is separable and of the form $\sum_{t=0}^{\infty} \beta^t \gamma^{-1} c_t^\gamma$. Then the marginal rate of substitution is $m_{t+1} = \beta(c_{t+1}/c_t)^{\gamma-1}$. Further assume that the consumption growth rate is bounded by

$$g_1 \leq \frac{c_{t+1}}{c_t} \leq g_2. \quad (43)$$

Then the marginal rate of substitution is bounded by

$$\beta g_2^{\gamma-1} \leq m_{t+1} \leq \beta g_1^{\gamma-1}, \quad (44)$$

and its standard deviation is bounded by

$$\text{std}(m) \leq \frac{\beta g_1^{\gamma-1} - \beta g_2^{\gamma-1}}{2}. \quad (45)$$

Combining inequalities (42) and (45), we obtain a lower bound on the RRA coefficient, $1 - \gamma$, as

$$\beta g_1^{\gamma-1} - \beta g_2^{\gamma-1} \geq \frac{2\{[E(R)/E(R_F)] - 1\}}{\text{std}(R)}. \quad (46)$$

Mehra and Prescott's parameter estimates are $E(R_F) = 1.01$ per year, $E(R) = 1.07$ per year, and $\text{std}(R) = .165$ per year. They assumed a two-state Markov process for the annual consumption growth rate. By the method of moments they estimated the annual consumption growth rate to be .982 or 1.054. Thus we can set $g_1 = .982$ and $g_2 = 1.054$. Our restriction (46) on the RRA coefficient becomes

$$.982^{\gamma-1} - 1.054^{\gamma-1} \geq \frac{.72}{\beta}. \quad (47)$$

For $\beta = .8$, the lower bound on the RRA coefficient $1 - \gamma$ is greater than or equal to 16; for $\beta = .9$, 14; for $\beta = 1$, 12; for $\beta = 1.1$, 11; and for $\beta = 1.2$, 10. The risk aversion is high, thereby illustrating the equity premium puzzle. Note that this conclusion is independent of the firm's leverage and of the correlation between consumption growth and the dividends on equity.

Essentially the lower bound on the consumption growth rate puts an upper bound on the marginal rate of substitution that is severe if utility is time separable. This causes the inability to explain the mean premium on equity returns. Rietz (1988) recognized the pivotal role of the lower bound on consumption growth. He proposed a model that allows for a disaster state, in which consumption may drop by as

much as 25 percent in one year. The model generates sufficient variability in the marginal rate of substitution and explains the observed mean premium on equity. Mehra and Prescott (1988) responded that the existence of such disasters has testable, but empirically unobserved, economic implications at times of impending disaster, such as the period of the Cuban Missile Crisis.

A plausible explanation of the puzzle, suggested by Mehra and Prescott, is that consumers are heterogeneous and the market is incomplete. Bewley (1982), Mankiw (1986), Scheinkman and Weiss (1986), and Scheinkman (1989) presented models with uninsurable risks. They illustrated that an econometrician may grossly overestimate risk aversion based on per capita consumption. The individuals' consumption growth rate may be substantially more variable than the per capita growth rate. Even with low variability of per capita consumption, the individuals' marginal rate of substitution may be sufficiently variable to explain the observed mean premium on equity.

Brainard and Summers (1987) and Kocherlakota (1988) allowed the equity to be levered, and Kocherlakota allowed the subjective discount rate to be negative ($\rho < 0$, i.e., $\beta > 1$). They found that the RRA coefficient must exceed 10 in order to generate the observed mean equity premium. This conclusion is consistent with the bounds on risk aversion derived in this section for time-separable utility.

Kocherlakota (1987) and Weil (1987) considered preferences that are not von Neumann–Morgenstern and found that the RRA coefficient must be high to explain the puzzle.

Nason (1988) generated the observed mean premium on equity by introducing state-nonseparable preferences in which the direct utility of consumption depends on past output. Whereas equilibrium consumption equals output in his model, the Euler equation and price paths are different from those implied by a direct utility function that depends on past consumption. One may view Nason's model as one in which utility exhibits habit persistence but the representative agent is myopic in that the agent disregards the effect of current consumption decisions on future utility.

The model in this paper generates the requisite high variability in the marginal rate of substitution in consumption with relatively low variability in the consumption growth rate through habit persistence in utility and low risk aversion. Essentially past consumption generates a subsistence level of consumption (which must be about 80 percent of the normal consumption rate in order to explain the mean equity premium, as in table 1). A small drop in consumption generates a large drop in consumption net of the subsistence level and a large drop in the marginal rate of substitution that makes it possible to match the observed equity premium with low risk aversion.

V. Concluding Remarks

One prediction of habit persistence is that the subsistence rate of consumption is positive. For the particular parameter values that explain the observed mean premium on equity, the subsistence rate of consumption is about 80 percent of the recent past consumption rate.

Habit persistence and durability of goods are opposing effects in that habit persistence tends to make certain lag coefficients in the Euler equation negative while durability tends to reverse their signs. Dunn and Singleton (1986), Eichenbaum, Hansen, and Singleton (1988), Gallant and Tauchen (1989), and Eichenbaum and Hansen (in press) used monthly data and estimated positive coefficients that are interpreted as evidence of durability. However, Ferson and Constantinides (1989) used quarterly and annual data and estimated negative coefficients that are interpreted as evidence of habit persistence with the subsistence level of the predicted order of magnitude. Furthermore, they rejected the time-separable model in favor of the model with habit persistence. Hansen and Jagannathan (1988) also found evidence in favor of habit persistence, using monthly data. Finally, Heaton (1988) examined the monthly and quarterly autocorrelations in consumption, while taking into account time aggregation, and interpreted his results as evidence of habit persistence.

Habit persistence departs from the familiar paradigm of state- and time-separable preferences. To become the new economic paradigm, habit persistence ought to be embedded in models of the business cycle, labor behavior, public finance, and so forth with preferences, technologies, and dynamics richer than the ones introduced in this paper and its predictions validated by empirical testing.

Appendix A

Proof of Theorem 1

a) The proof employs a technique that is applied in a different context by Davis and Norman (1987). Assume that an optimal policy $\{c(s), \alpha(s), t \leq s\}$ is given by $c(s) = c^*(s)$ and $\alpha(s) = \alpha^*(s)$, where $c^*(s)$ and $\alpha^*(s)$ are defined in equations (12) and (13). I shall prove that the optimal admissible policy for $0 \leq s \leq t$ is unique and is also given by $c(s) = c^*(s)$ and $\alpha(s) = \alpha^*(s)$.

b) For $s \geq t$, the capital increase is

$$dW(s) = \{[(\mu - r)\alpha^*(s) + r]W(s) - c^*(s)\}ds + \sigma\alpha^*(s)W(s)dw(s). \quad (A1)$$

Also,

$$dx(s) = [bc^*(s) - ax(s)]ds. \quad (A2)$$

Therefore,

$$\begin{aligned}
 d\left[W(s) - \frac{x(s)}{r+a-b}\right] &= \left\{[(\mu - r)\alpha^*(s) + r]W(s) - c^*(s) \right. \\
 &\quad \left. - \frac{bc^*(s) - ax(s)}{r+a-b}\right\}ds + \sigma\alpha^*(s)W(s)dw(s) \tag{A3} \\
 &= \left[W(s) - \frac{x(s)}{r+a-b}\right][nds + m\sigma dw(s)],
 \end{aligned}$$

which implies

$$d \ln \left[W(s) - \frac{x(s)}{r+a-b} \right] = \left(n - \frac{m^2\sigma^2}{2} \right) ds + m\sigma dw(s) \tag{A4}$$

with solution

$$\begin{aligned}
 W(s) - \frac{x(s)}{r+a-b} &= \left[W(t) - \frac{x(t)}{r+a-b} \right] \\
 &\quad \times \exp\left\{ \left(n - \frac{m^2\sigma^2}{2} \right) (s-t) + m\sigma[w(s) - w(t)] \right\}, \quad s \geq t. \tag{A5}
 \end{aligned}$$

For $s \geq t$,

$$\begin{aligned}
 &e^{-\rho(s-t)} E_t \{ [c^*(s) - x(s)]^\gamma \} \\
 &= h^\gamma \left[W(t) - \frac{x(t)}{r+a-b} \right]^\gamma \exp\left[-\rho(s-t) + \gamma \left(n - \frac{m^2\sigma^2}{2} \right) (s-t) \right] \\
 &\quad \times E_t \exp\{ \gamma m\sigma[w(s) - w(t)] \} \tag{A6}
 \end{aligned}$$

$$\begin{aligned}
 &= h^\gamma \left[W(t) - \frac{x(t)}{r+a-b} \right]^\gamma \exp\left\{ \left[-\rho + \gamma \left(n - \frac{m^2\sigma^2}{2} \right) + \frac{\gamma^2 m^2 \sigma^2}{2} \right] (s-t) \right\} \\
 &= h^\gamma \left[W(t) - \frac{x(t)}{r+a-b} \right]^\gamma \exp\left[-\frac{(r+a)h(s-t)}{r+a-b} \right]
 \end{aligned}$$

since

$$\begin{aligned}
 -\rho + \gamma \left(n - \frac{m^2\sigma^2}{2} \right) + \frac{\gamma^2 m^2 \sigma^2}{2} &= -\frac{1}{1-\gamma} \left[\rho - \gamma r - \frac{\gamma(\mu - r)^2}{2(1-\gamma)\sigma^2} \right] \\
 &= -\frac{(r+a)h}{r+a-b}. \tag{A7}
 \end{aligned}$$

Therefore,

$$\begin{aligned}
 V(W(t), x(t)) &= E_t \int_t^\infty e^{-\rho(s-t)} \gamma^{-1} [c^*(s) - x(s)]^\gamma ds \\
 &= \frac{h^\gamma}{\gamma} \left[W(t) - \frac{x(t)}{r+a-b} \right]^\gamma \int_t^\infty \exp \left[-\frac{(r+a)h(s-t)}{r+a-b} \right] ds \\
 &= \frac{(r+a-b)h^{\gamma-1}}{(r+a)\gamma} \left[W(t) - \frac{x(t)}{r+a-b} \right]^\gamma
 \end{aligned} \tag{A8}$$

since $(r+a)h/(r+a-b) > 0$. Note also that $V(W(t), x(t))$ is a C^2 function.

c) Define

$$M(t) \equiv \int_0^t e^{-\rho s} \gamma^{-1} [c(s) - x(s)]^\gamma ds + e^{-\rho t} V(W(t), x(t)) \tag{A9}$$

for an arbitrary policy $\{c(s), \alpha(s), 0 \leq s < t\}$. Applying Ito's lemma, we obtain $dM(t) = N(t)dt + e^{-\rho t} \alpha \sigma W V_w dw(t)$, where

$$\begin{aligned}
 N(t) &\equiv e^{-\rho t} \left(\gamma^{-1} (c-x)^\gamma - \rho V + \{[(\mu-r)\alpha + r]W - c\} V_w \right. \\
 &\quad \left. + \frac{\alpha^2 \sigma^2}{2} W^2 V_{ww} + (bc - ax) V_x \right).
 \end{aligned} \tag{A10}$$

Since $V_{ww} < 0$, $N(t)$ is concave in (c, α) . Suppressing momentarily the condition $0 \leq \alpha \leq 1$ and maximizing $N(t)$ with respect to (c, α) , we obtain the first-order conditions that are necessary and sufficient:

$$(c-x)^{\gamma-1} - V_w + bV_x = 0 \tag{A11}$$

and

$$(\mu-r)WV_w + \alpha \sigma^2 W^2 V_{ww} = 0. \tag{A12}$$

Solving, we obtain

$$c = x + h \left(W - \frac{x}{r+a-b} \right) = c^* \tag{A13}$$

and

$$\alpha = m \left(1 - \frac{x/W}{r+a-b} \right) = \alpha^*. \tag{A14}$$

Substituting $c^*(t)$ and $\alpha^*(t)$ in $N(t)$, we obtain $N(t) = 0$. Therefore,

$$\begin{aligned}
 dM(t) &\leq e^{-\rho t} \alpha \sigma W V_w dw(t) \quad \text{for arbitrary } (c, \alpha) \\
 &= e^{-\rho t} \alpha \sigma W V_w dw(t) \quad \text{for } (c = c^*, \alpha = \alpha^*),
 \end{aligned} \tag{A15}$$

and $M(t)$ is a supermartingale. Thus

$$E_0 \int_0^\infty e^{-\rho s} \gamma^{-1} [c(s) - x(s)]^\gamma ds = E_0 M(\infty) \leq EM(0) = V(W_0, x_0) \tag{A16}$$

with equality iff $c(s) = c^*(s)$ and $\alpha(s) = \alpha^*(s)$, for all s , $0 \leq s$. Therefore, the optimal policy for $0 \leq s < t$ is unique and is given by $c(s) = c^*(s)$ and $\alpha(s) = \alpha^*(s)$.

d) The policy $\{c^*(t), \alpha^*(t), t \geq 0\}$ obviously fulfills condition i of an admissible policy. It also fulfills condition ii. To see this, we proceed as in part b and derive (A5). Since $W_0 - [x_0/(r + a - b)] > 0$, it follows that $W(t) - [x(t)/(r + a - b)] > 0$. Since $h > 0$, it follows that $c(t) - x(t) > 0$. Since $x_0 > 0$, then $c(0) > 0$, $x(t) > 0$, and $c(t) > 0$. Finally, $\int_0^t c(s) < \infty$, for all t almost surely.

The policy also fulfills condition iii of an admissible policy. Since $0 \leq m \leq 1$ and $0 < 1 - \{[x(t)/W(t)]/(r + a - b)\} < 1$, it follows that $0 \leq \alpha^*(t) < 1$.

Finally, $W(t) > 0$ since $W(t) - [x(t)/(r + a - b)] > 0$. This fulfills the last condition of an admissible policy.

e) Under the optimal policy, equation (A8) gives the derived utility of capital as in equation (15). Equation (A5) gives the capital as in equation (16).

f) To find the consumption growth rate, we can use (A13) and (A5) to write

$$\begin{aligned} \ln(c - x) &= \ln h + \ln\left(W - \frac{x}{r + a - b}\right) \\ &= \ln h + \ln\left(W_0 - \frac{x_0}{r + a - b}\right) + \left(n - \frac{m^2\sigma^2}{2}\right)t + m\sigma w(t), \end{aligned} \tag{A17}$$

$$\frac{dc - dx}{c - x} = ndt + m\sigma dw(t), \tag{A18}$$

$$\begin{aligned} \frac{dc}{c} &= \frac{1}{c} [dx + (c - x)ndt + (c - x)m\sigma dw(t)] \\ &= \frac{1}{c} [bc - ax + (c - x)n]dt + \left(1 - \frac{x}{c}\right)m\sigma dw(t) \end{aligned} \tag{A19}$$

$$= \left[n + b - \frac{(n + a)x}{c}\right]dt + \left(1 - \frac{x}{c}\right)m\sigma dw(t),$$

which proves (17).

Appendix B

Proof of Theorem 2

i) We can combine equations (3) and (17) and obtain the diffusion equation for z as

$$dz = [b - (n + a - m^2\sigma^2)z - m^2\sigma^2z^2](1 - z)dt - z(1 - z)m\sigma dw(t). \tag{B1}$$

From (B1) we obtain the diffusion equation for $y = z/(1 - z)$ as

$$dy = [b - (n + a - b - m^2\sigma^2)y]dt - m\sigma y dw(t). \tag{B2}$$

The forward, or Fokker-Planck, equation for the density $p_y(y(t); y_0, t)$, $0 < t < \infty$, is

$$\frac{1}{2} \frac{\partial^2}{\partial y^2} (m^2\sigma^2y^2p_y) - \frac{\partial}{\partial y} \{[b - (n + a - b - m^2\sigma^2)y]p_y\} = \frac{\partial p_y}{\partial t}. \tag{B3}$$

This forward equation is a member of the class of forward equations studied by Wong (1964). Specializing the results of Wong for the problem at hand, we conclude that, for $0 \leq y < \infty$, the stationary distribution of $p_y(y)$ exists and is the solution of the Pearson equation

$$\frac{1}{2} \frac{\partial}{\partial y} (m^2 \sigma^2 y^2 p_y) - [b - (n + a - b - m^2 \sigma^2) y] p_y = 0 \tag{B4}$$

subject to the normalization

$$\int_0^\infty p_y(y) dy = 1. \tag{B5}$$

We rewrite equation (B4) as

$$\frac{1}{2} m^2 \sigma^2 y^2 \frac{\partial}{\partial y} p_y - [b - (n + a - b) y] p_y = 0. \tag{B6}$$

The solution of equation (B6) is equation (20), and the normalizing constant k is stated in terms of a gamma function in equation (21).

A mode, \hat{y} , of the stationary distribution satisfies $\partial p_y / \partial y = 0$, and we obtain $b - (n + a - b) \hat{y} = 0$ with unique solution given by equation (22). We integrate equation (B6) by parts and obtain

$$\frac{1}{2} m^2 \sigma^2 [y^2 p_y]_0^\infty - m^2 \sigma^2 \int_0^\infty y p_y dy - b + (n + a - b) \int_0^\infty y p_y dy = 0. \tag{B7}$$

Inspection of equation (20) implies $y^2 p_y \rightarrow 0$ as $y \rightarrow 0$. Also, condition (19) and equation (20) imply $y^2 p_y \rightarrow 0$ as $y \rightarrow \infty$. Then equation (B7) gives the mean value of y as in equation (23). Condition (19) guarantees that the mean is finite.

ii) Since $z = y/(1 + y)$ and $0 \leq y < \infty$, it follows that z is monotone increasing in y in the domain of y and $0 \leq z < 1$. Since $y(t)$ has a stationary distribution, so does $z(t)$. The density of the stationary distribution of z is $p_z(z)$:

$$p_z(z) = p_y(y) \frac{dy}{dz} = (1 - z)^{-2} p_y\left(\frac{z}{1 - z}\right). \tag{B8}$$

Combining equations (B8) and (20), we obtain equation (24).

If \hat{z} is a mode of the stationary distribution, it must satisfy $\partial p_z(z) / \partial z = 0$, which, on simplification, becomes

$$m^2 \sigma^2 \hat{z}^2 - (n + a) \hat{z} + b = 0. \tag{B9}$$

The left-hand side of (B9) equals $b > 0$ at $\hat{z} = 0$ and $m^2 \sigma^2 - n - a + b < 0$ at $\hat{z} = 1$. Therefore, the quadratic has only one root in the range $0 \leq \hat{z} \leq 1$. The stationary distribution has a single mode at \hat{z} given by equation (25). This completes the proof.

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